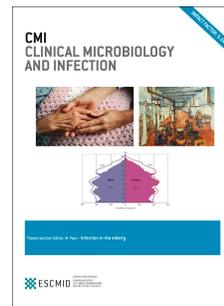


# Journal Pre-proof

Predicting critical illness on initial diagnosis of COVID-19 based on easily-obtained clinical variables: development and validation of the PRIORITY model

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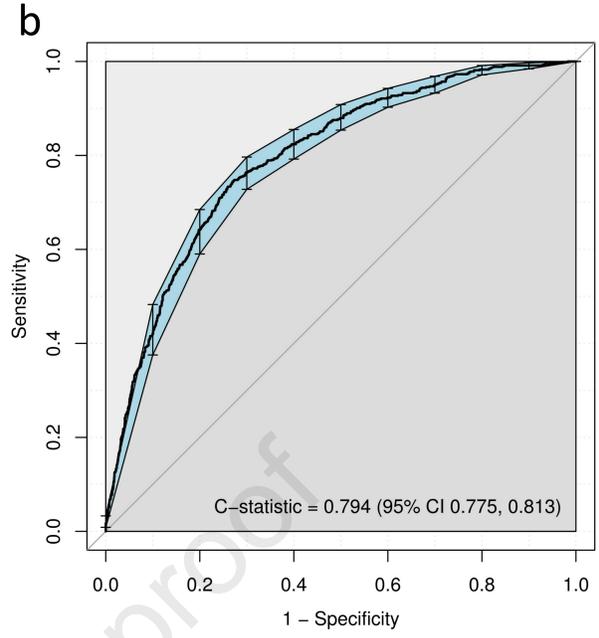
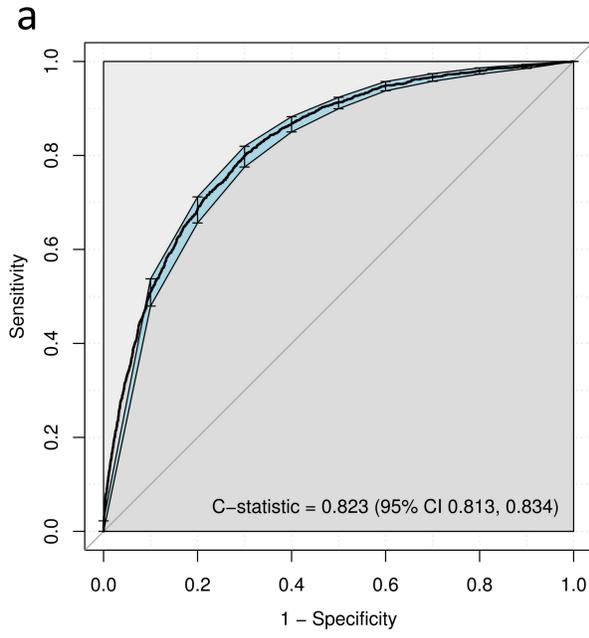
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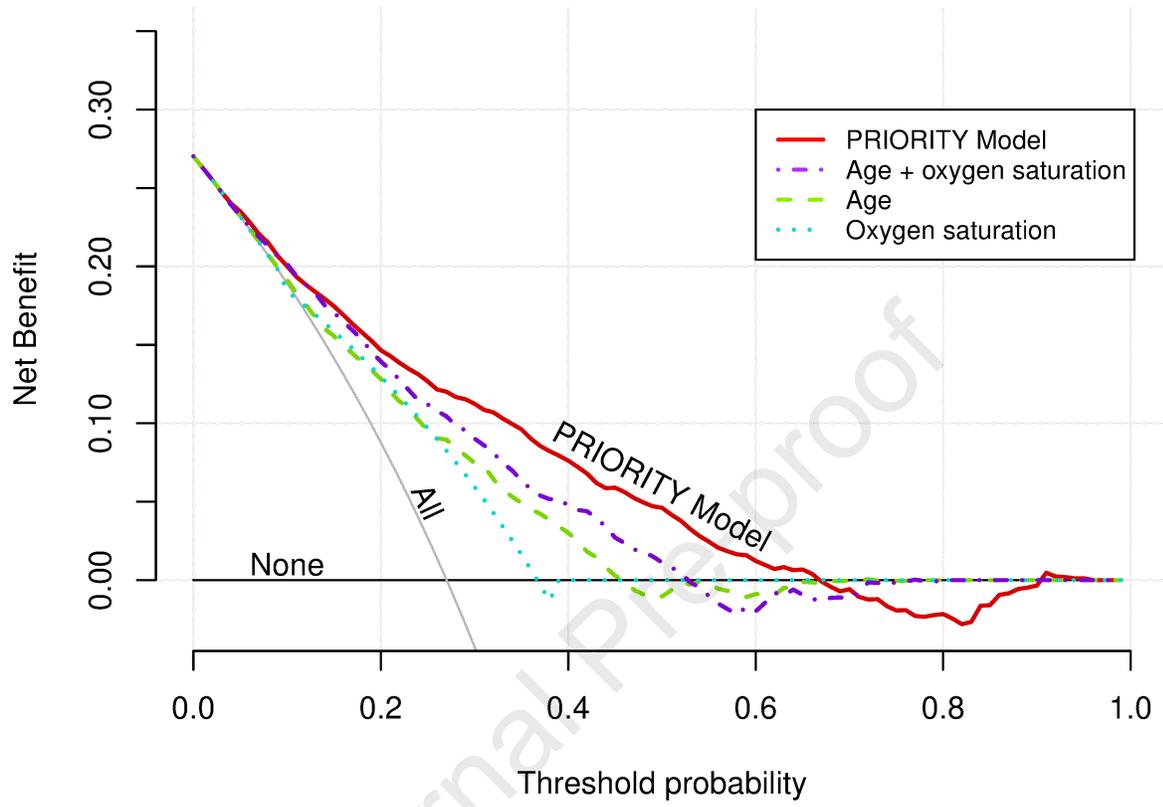
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1 **TITLE**

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 3 **variables: Development and validation of the PRIORITY model**

4  
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 45  
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**50 Objectives**

51 We aimed to develop and validate a prediction model, based on clinical history and examination findings  
52 on initial diagnosis of COVID-19, to identify patients at risk of critical outcomes.

**53 Methods**

54 We used data from the SEMI-COVID-19 Registry, a cohort of consecutive patients hospitalized for  
55 COVID-19 from 132 centers in Spain (23 March to 21 May, 2020). For the development cohort tertiary  
56 referral hospitals were selected, while the validation cohort included smaller hospitals. The primary  
57 outcome was a composite of in-hospital death, mechanical ventilation or admission to intensive care  
58 unit. Clinical signs and symptoms, demographics, and medical history ascertained at presentation were  
59 screened using least absolute shrinkage and selection operator, and logistic regression was used to  
60 construct the predictive model.

**61 Results**

62 There were 10,433 patients, 7,850 in the development cohort (primary outcome 25.1%, 1,967/7,850)  
63 and 2,583 in the validation cohort (outcome 27.0%, 698/2,583). The PRIORITY model included: age,  
64 cardiovascular disease, chronic kidney disease, dyspnea, tachypnea, confusion, systolic blood pressure,  
65 and  $SpO_2 \leq 93\%$  or oxygen requirement. The model showed high discrimination for critical illness in both  
66 the development (C-statistic 0.823; 95% confidence interval [CI] 0.813, 0.834) and validation (C-statistic  
67 0.794; 95% CI 0.775, 0.813) cohorts. A freely available web-based calculator was developed based on  
68 this model (<https://www.evidencio.com/models/show/2344>).

**69 Conclusions**

70 The PRIORITY model, based on easily-obtained clinical information, had good discrimination and  
71 generalizability for identifying COVID-19 patients at risk of critical outcomes.

72

## 73 INTRODUCTION

74 Coronavirus disease 2019 (COVID-19) has spread globally, with a clinical spectrum ranging from an  
75 asymptomatic state to critical illness [1-3]. Notably, Spain was one of the countries with the highest  
76 incidence of COVID-19 during the first pandemic peak [4]. To optimize the use of limited healthcare  
77 resources, it would be essential to identify, as early as possible, those patients who are at high risk of  
78 progressing to critical illness.

79 To date, studies of COVID-19 prognostic factors have focused on laboratory measurements and  
80 radiological examinations obtained following admission [5-15], which are not available in outpatient or  
81 resource-limited settings. Recently published well-developed models tend not to include clinical  
82 variables obtained from history and examination carried out on initial assessment [9-13]. Where one  
83 machine learning model has addressed basic clinical features, it has narrowed down the prediction to  
84 the mortality outcome only and lacks wider generalizability [16]. Furthermore, a critical appraisal of the  
85 COVID-19 models has shown poor reporting and high risk of bias [14].

86 Prediction models based on easy-to-collect data have previously been developed for other infectious  
87 diseases, e.g. meningitis and pneumonia [17-19]. As a global health emergency, management of COVID-  
88 19 would benefit from a prediction model that could be readily applied for initial diagnosis. Therefore,  
89 we developed and externally validated a prediction model, based on easily obtained clinical measures at  
90 presentation with confirmed COVID-19 diagnosis, to identify patients at risk of developing critical  
91 outcomes.

## 92 METHODS

### 93 Study design and data source

94 This study was based on the SEMI (Sociedad Española de Medicina Interna) COVID-19 Registry [20]. It is  
95 an ongoing multicenter nationwide cohort of consecutive patients hospitalized for COVID-19 across  
96 Spain. Eligibility criteria were age  $\geq 18$  years, confirmed diagnosis of COVID-19, defined as a positive  
97 result on real-time reverse-transcription-polymerase-chain-reaction (RT-PCR) for the presence of SARS-  
98 CoV-2 in nasopharyngeal swab specimens or sputum samples, first hospital admission for COVID-19, and  
99 hospital discharge or in-hospital death [20]. The SEMI-COVID-19 Registry was approved by the Provincial  
100 Research Ethics Committee of Málaga (Spain) and the Institutional Research Ethics Committees of each  
101 participating hospital.

102 For the study, we retrieved from the Registry clinical baseline data, history of previous medication, and  
103 comorbidities collected on admission, as well as complications during hospitalization and status at  
104 discharge. We used data from patients admitted in 132 hospitals between March 23 and May 21, 2020.  
105 We chose hospital complexity as the criterion to assess the transportability of the prognostic model in a  
106 setting other than the one in which it was derived [21, 22]. Patients admitted to tertiary referral  
107 hospitals ( $\geq 300$  beds, according to the Ministry of Health of Spain [23]) were selected for the  
108 development cohort, while patients from smaller hospitals ( $< 300$  beds) were included in a separate  
109 validation cohort.

#### 110 **Outcome description**

111 The primary outcome, critical illness during hospitalization, was defined as the composite of in-hospital  
112 death, mechanical ventilation or admission to intensive care unit (ICU), according to previously  
113 published studies [10, 24-25].

#### 114 **Potential predictors**

115 To develop a predictive model based only on easily measurable variables registered at admission, we

116 considered clinical signs and symptoms, demographic variables, and medical history. An initial list of 29  
117 candidate variables was selected based on review of the existing evidence [5-16], clinical plausibility and  
118 relevance to clinical care. To improve consensus on model applicability, a 1-round online questionnaire  
119 was conducted among a multidisciplinary panel of 24 physicians involved in COVID-19 clinical  
120 management at nursing homes, emergency departments, primary care centers and hospitalization  
121 wards (6 per each setting). The panelists were asked to rate (on a 9-point Likert scale) the  
122 availability/reliability of each predictor, its ability to predict the outcome, the best way to merge  
123 predictors of rare occurrence, and the maximum number of variables the model should contain.  
124 Agreement was considered when  $\leq 7$  panelists rated outside the 3-point region containing the median  
125 [26].

#### 126 **Statistical analysis**

127 The predictive model, called PRIORITY, was presented as the formula for estimating the probability of  
128 COVID-19 critical illness outcome, as well as an associated web-based calculator. Patients'  
129 characteristics were summarized as frequencies and percentages or mean and standard deviation.  
130 Statistical analysis was performed using R v4.0.0, with mice, mfp, glmnet, pROC, rms and rmda  
131 packages.

132 *Model development:* Missing values in the potential predictors were imputed using single imputation, a  
133 reasonable alternative to multiple imputation when dealing with relatively few missings [27]. A  
134 stochastic single imputation dataset was created for both cohorts (development and validation) through  
135 multiple imputation by chained equations. Quantitative variables were kept as continuous to avoid loss  
136 of prognostic information, and non-linear relationships were modelled by multivariate fractional  
137 polynomials with a maximum of 2 degrees of freedom [28]. The least absolute shrinkage and selection  
138 operator (LASSO) was the feature selection method used to reduce the number of predictors down to

139 the maximum agreed by the expert panel [29]. Briefly, the potential predictors were entered into the  
140 LASSO regularization process, which penalizes the coefficients by gradually shrinking them to zero. We  
141 selected the penalty parameter ( $\lambda$ ) that minimized the deviance within the given maximum number of  
142 predictors. Those variables with non-zero coefficients were retained for risk estimation using a logistic  
143 regression model. Coefficients were presented as odds ratios (OR) and 95% confidence intervals (CI).

144 *Model performance:* Nagelkerke's  $R^2$  and Brier score were used as overall performance measures. We  
145 assessed the discriminative ability of the model using the C-statistic, calculated as the area under the  
146 Receiver Operating Characteristic curve, with 95% CI. Calibration of the model was visually assessed by  
147 plotting deciles of predicted vs. observed probabilities, and the calibration slope with 95% CI was  
148 calculated [22].

149 *Model validation:* Internal validation was carried out to assess optimism-corrected performance by  
150 repeating the entire model development over 1,000 bootstrap samples drawn from the development  
151 cohort [27]. We externally validated the model in a separate cohort of patients admitted at less-complex  
152 hospitals to evaluate model transportability [21]. Within this validation cohort, we reassessed model  
153 performance and compared its discrimination ability with models based on oxygen saturation and/or  
154 age, the most discriminating univariate predictors for in-hospital mortality previously reported [15]. We  
155 also undertook a decision curve analysis, a method to ascertain the adequacy of prediction models  
156 based on the relative value of benefits (true positives) and harms (false positives) [30]. We plotted the  
157 net benefit of the models for the full range of critical illness probability thresholds.

158 *Sensitivity analysis:* To assess the impact of assumptions adopted in the model development, we carried  
159 out a complete-case analysis, using only those patients with complete data in the potential predictors.  
160 We also developed models without restricting the maximum number of predictors ( $\lambda$  at one-standard-  
161 error of the minimum) or using linear continuous predictors instead of non-linear terms.

162 **RESULTS**

163 We considered data from 10,433 patients included in the SEMI-COVID-19 Registry. The development  
164 cohort included 7,850 patients, of which 1,967 (25.1%) presented critical illness (650 [8.3%] admitted to  
165 the ICU and 1,598 [20.4%] died). The mean age was  $65.8 \pm 16.4$  years and 57.2% (4,483/7,834) were  
166 male, ageusia/anosmia, asthenia/anorexia, headache, gastrointestinal symptoms were excluded.  
167 Consensus was achieved for including a range between 5 and 9 variables in the final model. For  
168 transparency, univariate analysis is shown on *Supplementary Table S1*, even though it was not part of  
169 the model development process. The 21 potential predictors were included in the LASSO selection  
170 process, retaining a subset of 9 variables as the best predictors of critical illness (*Supplementary Figure*  
171 *S1*). A multivariable logistic regression model was then fitted with these 9 variables. All of them, except  
172 for moderate or severe dependency, were statistically significant (*Table 2*).

173 Based on the logistic regression model, the probability of critical COVID-19 illness could be calculated as:

174 Probability (%) =  $100/(1 + \exp(-z))$ , where  $z = -4.665 + 2.663 \cdot [(Age/100)^2] + 0.164 \cdot [Dependency] +$   
175  $0.316 \cdot [Cardiovascular \ disease] + 0.586 \cdot [Chronic \ kidney \ disease] + 0.504 \cdot [Dyspnea] +$   
176  $0.844 \cdot [1/(SBP/100)^2] + 0.911 \cdot [Tachypnea] + 1.200 \cdot [SpO_2 \leq 93\% \text{ or oxygen requirement}] +$   
177  $0.681 \cdot [Confusion]$ .

178 All predictors were coded as binary variables (1 when present and 0 when absent) except for age (years)  
179 and systolic blood pressure (SBP, mmHg). We also developed an online calculator based on this model  
180 (*Supplementary Figure S2*), accessible at <https://www.evidencio.com/models/show/2344>.

181 In the development cohort, the PRIORITY model had an  $R^2$  of 0.347 and a Brier score of 0.138. The  
182 apparent C-statistic was 0.823 (95% CI 0.813, 0.834) (*Figure 1a*). After bootstrap internal validation,  
183 optimism-corrected C-statistic was 0.821 (95% CI 0.810, 0.832). The model showed good calibration

184 across the range of predicted probabilities within the development cohort (calibration slope 0.996, 95%  
185 CI 0.989, 0.999; *Supplementary Figure S3a*).

#### 186 **External validation**

187 The validation cohort included 2,583 patients, of which 698 (27.0%) presented critical illness (200 [7.7%]  
188 admitted to the ICU and 594 [23.0%] died). The mean age was  $69.5 \pm 16.0$  years, 54.8% (1,415/2,580)  
189 were male (*Table 1*). The PRIORITY model showed good discrimination for critical illness within the  
190 validation cohort (C-statistic 0.794, 95% CI 0.775, 0.813) (*Figure 1b*), and a calibration slope of 0.875,  
191 95% CI 0.790, 0.960 (*Supplementary Figure S3b*).

192 Our model compared well against the risk stratification based on univariate models including age (C-  
193 statistic 0.707, 95% CI 0.686, 0.729) or SpO<sub>2</sub>≤93%/oxygen requirement at admission (C-statistic 0.652,  
194 95% CI 0.635, 0.670) as sole predictors. Likewise, the PRIORITY model had better discrimination ability  
195 than the model including both age and SpO<sub>2</sub>≤93%/oxygen supply (C-statistic 0.751, 95% CI 0.731, 0.771).

196 Additionally, decision curve analysis showed that the PRIORITY model had higher net benefit across a  
197 wide range of threshold probabilities for developing critical illness compared to risk stratification using  
198 age and/or SpO<sub>2</sub>≤93%/oxygen supply (*Figure 2*).

#### 199 **Sensitivity analysis**

200 We carried out a complete-case analysis selecting as development cohort the 5,513 patients with  
201 complete data on the 21 potential predictors and the outcome. The resulting model had the same  
202 predictors as the PRIORITY model with apparent C-statistic 0.813 (95% CI 0.800, 0.826) and calibration  
203 slope 0.993 (95% CI 0.986, 0.997). Next, we fitted a new model with no restriction in maximum number  
204 of variables, resulting in a model which added sex, diabetes mellitus, malignancy, immunocompromised

205 status, pulmonary rates, and heart rate cubed to the predictors in the PRIORITY model. C-statistic was  
206 0.831 (95% CI 0.821, 0.842) and slope 0.990 (95% CI 0.986, 0.996). Likewise, we fitted an alternative  
207 model using linear continuous predictors instead of non-linear terms, which included sex but excluded  
208 the systolic blood pressure. C-statistic was 0.823 (95% CI 0.812, 0.833) and slope 0.994 (95% CI 0.988,  
209 0.999).

## 210 **DISCUSSION**

211 We developed and validated a new clinical risk model to predict COVID-19 critical illness based on nine  
212 simple clinical features easily available on initial assessment, which would be useful in resource-limited  
213 or out-of-hospital settings without access to other complementary tests. The model was well calibrated,  
214 had good discrimination, and performed robustly in an external validation cohort. Moreover, it showed  
215 a potential clinical benefit in a variety of scenarios covering different healthcare situations over a range  
216 of threshold probabilities for critical illness. The web-based calculator can facilitate its immediate  
217 application for frontline clinicians.

218 Previously, an external validation of 22 prognostic models showed that none of the multivariate models  
219 offered incremental value for patient stratification compared to oxygen saturation or age [15]. In this  
220 regard, the PRIORITY model showed higher discriminative ability and net benefit than age and/or oxygen  
221 saturation. Additionally, despite its simplicity, our model had a similar performance to previously  
222 published prognostic tools including laboratory and imaging tests [9-16].

223 It is worth noting that the PRIORITY model could be applied in triage, using easily measurable variables  
224 available in settings without access to laboratory or radiology tests, identifying high-risk patients for  
225 referral to hospital. This model could be useful for supporting clinical management decisions over a  
226 range of risk thresholds for critical illness which could be considered as relevant in clinical practice. The

227 choice of thresholds will vary across different regions, according to changing epidemiological situations  
228 and availability of health resources. For example, under pandemic peak pressure or low-resource  
229 healthcare systems, policy-makers may consider a cut-off point up to 20%, a threshold that will be  
230 associated with higher reduction in unnecessary critical care admissions. However, at low risk of  
231 overwhelming the critical care capacity, a lower threshold may be considered at the expense of  
232 unnecessary referrals. We recommend objectively defining specific cut-off points considering the  
233 circumstances and the availability of health resources.

234 This study has several methodological strengths maximizing internal and external validity [23]. To the  
235 best of our knowledge, this is the first generalizable COVID-19 predictive model built with simple clinical  
236 information excluding imaging and laboratory data. We developed and validated the model in a large  
237 multicenter, national cohort. The methodology was rigorous, avoiding data-driven predictor selection  
238 and biases that affected previous studies [14]. The practical application of the model was maximized by  
239 forging an agreement among an expert panel on key issues. Moreover, the model was validated in a  
240 separate cohort of patients admitted in smaller hospitals, showing transportability to a setting with a  
241 different level of healthcare [21, 22].

242 The strength of our findings should be interpreted in light of some limitations. First, although we  
243 carefully selected easily available clinical and demographic variables, the data were collected at the time  
244 of hospital admission, which represents an important selection bias that would require further studies in  
245 an outpatient setting. Second, it could be suggested that, taking into account the situation of healthcare  
246 pressure, data quality may be affected. In this regard, it is notable that in this study missing data were  
247 relatively low and we used imputation to reduce their impact. Third, since the COVID-19 pandemic has  
248 demonstrated significant differences between countries and time periods, it could affect the  
249 applicability of the model to other settings. However, we considered this early pandemic period in Spain

250 to reflect a scenario with an overwhelmed healthcare system, where our predictive model could be  
251 particularly useful. Nevertheless, further studies introducing factors such as viral strains, healthcare  
252 system actions, new treatments, or vaccination, could improve the applicability of the PRIORITY model.  
253 Lastly, even though we compared the net benefit of using the model with discrimination based on  
254 oxygen saturation and/or age, its real clinical usefulness would require comparison with the best  
255 existing scores or the clinician's decision.

256 In summary, we developed and validated a new prediction model, called PRIORITY, to estimate the risk  
257 of critical illness in patients with COVID-19 based on nine clinical variables easily measurable in  
258 resource-limited or out-of-hospital settings. The study could provide underpinning evidence to inform  
259 decision-making in health systems under pandemic pressure.

260

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274 data, logistic and administrative support. The authors declare that there are no conflicts of interest.

#### 275 **AUTHORS CONTRIBUTIONS**

276 MML, LAVN and MFF planned, conceived the study, analysed and interpreted the data. MML, LAVN,  
277 MFF and LM wrote the original draft of the manuscript. MRR, SLG, FAF, JLBP, JAVN, ECM, ACEA, SJFC,  
278 JLA, PMPF, AP, AMAS, ASA, BGL, JLP, JSC, PCP, GMGG, JMNC, JMCR and RGH contributed to read and  
279 approved the final version of the manuscript. MML and LAVN are joint first authors. The corresponding  
280 author attests that all listed authors meet authorship criteria and that no others meeting the criteria  
281 have been omitted. LM is the guarantor.

#### 282 **TRANSPARENCY DECLARATION**

283 The correspondent author affirms that the manuscript is an honest, accurate, and transparent account  
284 of the study being reported; that no important aspects of the study have been omitted; and that any  
285 discrepancies from the study as planned have been explained.

#### 286 **DISSEMINATION TO PARTICIPANTS AND RELATED PATIENT AND PUBLIC COMMUNITIES**

287 Results of this study have been made available to the public through an open access preprint posted to  
288 MedRxiv (doi: 10.1101/2020.11.27.20237966). The Spanish Society of Internal Medicine (SEMI) shares  
289 the results of the studies derived from the SEMI-COVID-19 Registry through its public facing website  
290 (<https://www.fesemi.org/investigacion/proyectos/registro-semi-covid-19>) and its twitter account  
291 (@Sociedad\_SEMI).

#### 292 **DATA SHARING**

293 The data that support the findings of this study are available on request from the SEMI-COVID-19

294 Scientific Committee and the Registry Coordinating Center.

295

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297 This study did not receive funding.

298 **DISCLOSURES**

299 The authors declare no conflict of interest.

300

301

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## 378 TABLES

**Table 1. Demographic and clinical characteristics among patients included in the development and validation cohorts.**

		Development cohort		Validation cohort	
		No of patients (%) or mean $\pm$ SD	Total No (%)	No of patients (%) or mean $\pm$ SD	Total No (%)
<b>Characteristics of the population</b>					
Critical illness		1967 (25.1%)	7850 (100%)	698 (27.0%)	2583 (100%)
Age [years]		65.8 $\pm$ 16.4	7816 (99.6%)	69.5 $\pm$ 16.0	2575 (97.3%)
Male		4483 (57.2%)	7834 (99.8%)	1415 (54.8%)	2580 (99.9%)
Ethnicity	Caucasian	6836 (89.1%)	7677 (98.8%)	2340 (91.0%)	2572 (99.6%)
	Latino	693 (9.0%)		193 (7.5%)	
	Other	148 (1.9%)		39 (1.5%)	
Smoking history	Never	5270 (70.9%)	7433 (94.7%)	1625 (65.7%)	2475 (95.8%)
	Former smoker	1764 (23.7%)		718 (29.0%)	
	Active Smoker	399 (5.4%)		139 (5.3%)	
<b>Medical history</b>					
Obesity		1665 (23.7%)	7012 (89.3%)	584 (24.3%)	2401 (93.0%)
Hypertension		3803 (48.6%)	7833 (99.8%)	1444 (56.1%)	2576 (99.7%)
Diabetes mellitus		1440 (18.4%)	7820 (99.6%)	509 (19.8%)	2570 (99.5%)
Cardiovascular disease		1974 (25.3%)	7800 (99.4%)	806 (31.7%)	2545 (98.5%)
Pulmonary diseases		1625 (20.9%)	7776 (99.1%)	576 (22.6%)	2583 (98.9%)
Severe chronic kidney disease		488 (6.2%)	7825 (99.7%)	163 (6.3%)	2583 (99.7%)
Malignancy		793 (10.2%)	7803 (99.4%)	259 (10.1%)	2571 (99.5%)
Immunocompromised status		650 (8.6%)	7549 (96.2%)	187 (7.6%)	2473 (95.7%)
Dependency (moderate/severe)		1129 (14.7%)	7701 (98.1%)	605 (23.7%)	2555 (98.9%)
<b>Symptoms at admission</b>					
Fever		5138 (67.0%)	7663 (97.6%)	1670 (65.6%)	2544 (98.5%)
Dyspnea		4427 (56.7%)	7805 (99.4%)	1523 (59.4%)	2562 (99.2%)
<b>Clinical signs and physical exploration at admission</b>					
SBP [mmHg]		129.0 $\pm$ 21.5	7430 (94.6%)	127.6 $\pm$ 21.0	2451 (94.9%)
HR [beats/minute]		88.6 $\pm$ 17.4	7500 (95.5%)	87.5 $\pm$ 17.5	2504 (96.9%)
Tachypnea (> 20 breaths/min)		2271 (29.9%)	7604 (96.9%)	879 (35.1%)	2504 (96.9%)
SpO <sub>2</sub> $\leq$ 93% or oxygen requirement at presentation		4152 (52.9%)	7842 (99.9%)	1605 (62.1%)	2583 (100%)
Pulmonary rales		4630 (60.7%)	7626 (97.1%)	1588 (63.6%)	2495 (96.6%)
Confusion		849 (11.0%)	7736 (98.5%)	384 (15.1%)	2546 (98.6%)

379 SD: standard deviation.

380 Obesity is defined as Medical history or body mass index  $\geq 30$  kg/m<sup>2</sup>.  
 381 Cardiovascular disease: history of cerebrovascular disease, peripheral arterial disease, myocardial infarction, angina pectoris, heart failure or  
 382 atrial fibrillation.  
 383 Pulmonary diseases: chronic obstructive pulmonary disease, obstructive sleep apnea/hypopnea syndrome and asthma.  
 384 Severe chronic kidney disease: History of serum creatinine level  $> 3$  mg/dl or history of dialysis.  
 385 Malignancy: History of solid tumor, leukemia or lymphoma.  
 386 Immunocompromised status: History of autoimmune diseases, solid organ transplant recipients, HIV infection or previous immunosuppressive  
 387 treatment including systemic steroids.  
 388 Dependency (moderate/severe): moderate or severe dependency for activities of daily living (Barthel index score  $<60$ ). Fever: Temperature  $\geq$   
 389  $38^{\circ}\text{C}$  or history of fever.  
 390 HR: Heart rate.  
 391 SBP: Systolic blood pressure.  
 392 SpO<sub>2</sub>: Peripheral oxygen saturation.

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**Table 2. Multivariate logistic regression of critical illness prediction in COVID-19.**

Predictors	Odds ratio	95% CI
(Age/100) <sup>2</sup> [Age in years]*	14.339	10.054, 20.532
Cardiovascular disease	1.372	1.195, 1.573
Severe chronic kidney disease	1.797	1.433, 2.252
Dyspnea	1.655	1.451, 1.891
1/(SBP/100) <sup>2</sup> [SBP in mmHg]*	2.326	1.837, 2.951
Tachypnea ( $>20$ breaths/min)	2.487	2.192, 2.824
SpO <sub>2</sub> $\leq 93\%$ or oxygen requirement	3.320	2.889, 3.819
Confusion	1.976	1.642, 2.380
Dependency (Moderate or severe)	1.178	0.989, 1.404

394 Predictors in the PRIORITY model retained after LASSO feature selection. Model coefficients were derived from a multivariate logistic  
 395 regression, and presented as odds ratios (OR) and 95% confidence intervals (95% CI).

396 Variables entered into the LASSO feature selection process were: age as a squared term, sex, ethnicity, smoking history, obesity, hypertension,  
 397 diabetes mellitus, cardiovascular disease, pulmonary diseases, severe chronic kidney disease, malignancy, immunocompromised status,  
 398 dependency, fever, dyspnea, systolic blood pressure (SBP) as the inverse of a quadratic term, heart rate (HR) as a cubic term, tachypnea,  
 399 peripheral oxygen saturation (SpO<sub>2</sub>)  $\leq 93\%$  on room air or oxygen requirement at presentation, pulmonary rates, and confusion. All predictors  
 400 were coded as binary variables (1 when present and 0 when absent) except for age (years), SBP (mmHg) and HR (bpm).

401 \* Continuous predictors modelled as fractional polynomial terms, including rescaling when the range of values of the predictor was reasonably  
 402 large. As interpretability of the effect of non-linear continuous predictors can be difficult, linear local approximations of ORs for 10-unit  
 403 variations are provided at selected values.

404 ORs for age (10-year increments): OR (50/40) = 1.271; OR (70/60) = 1.414; OR (90/80) = 1.573.

405 ORs for SBP (10 mmHg decreases): OR (110/120) = 1.118; OR (90/100) = 1.219; OR (70/80) = 1.497.

406 Approximated ORs are provided for illustrative purposes only and were not used for making predictions.  
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## 410 FIGURE LEGENDS

411 Figure 1. Discriminatory ability of the PRIORITY model in the development (a) and validation (b) cohorts.  
 412 Discriminative ability was assessed using the C-statistic, as the area under the Receiver Operating  
 413 Characteristic curve, with 95% confidence intervals (CI) computed with 1,000 bootstrap replicates.

414

415 Figure 2. Decision curve analysis within the validation cohort. Clinical usefulness of the PRIORITY model  
416 compared to risk stratification based on oxygen saturation (binary:  $SpO_2 \leq 93\%$  or oxygen requirement)  
417 and/or age (quadratic term). The x-axis represents the whole range of decision threshold probabilities  
418 for critical illness ( $p_t$ ) and the y-axis the net benefit (NB). NB calculated as: True positives/N – (False  
419 positives/N)\*( $p_t/(1-p_t)$ ), with N total sample size.

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